## A Search for $B^{+} \rightarrow \tau^{+} \boldsymbol{\nu}$ decays with hadronic $B$ tags

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We present a search for the decay $B^{+} \rightarrow \tau^{+} \nu$ using $383 \times 10^{6} B \bar{B}$ pairs collected at the $\Upsilon(4 S)$ resonance with the $B A B A R$ detector at the SLAC PEP-II $B$ Factory. We select a sample of events with one completely reconstructed $\operatorname{tag} B$ in a hadronic decay mode $\left(B^{-} \rightarrow D^{(*) 0} X^{-}\right)$, and examine the rest of the event to search for a $B^{+} \rightarrow \tau^{+} \nu$ decay. We identify the $\tau$ lepton in the following modes: $\tau^{+} \rightarrow e^{+} \nu \bar{\nu}, \tau^{+} \rightarrow \mu^{+} \nu \bar{\nu}, \tau^{+} \rightarrow \pi^{+} \bar{\nu}$ and $\tau^{+} \rightarrow \pi^{+} \pi^{0} \bar{\nu}$. We find a $2.2 \sigma$ excess in data and measure a branching fraction of $\mathcal{B}\left(B^{+} \rightarrow \tau^{+} \nu\right)=\left(1.8_{-0.8}^{+0.9}\right.$ (stat.) $\pm 0.4$ (bkg. syst.) $\pm$ 0.2 (other syst.)) $\times 10^{-4}$. We calculate the product of the $B$ meson decay constant $f_{B}$ and $\left|V_{u b}\right|$ to be $f_{B} \cdot\left|V_{u b}\right|=\left(10.1_{-2.5}^{+2.3}(\text { stat. })_{-1.5}^{+1.2}(\right.$ syst. $\left.)\right) \times 10^{-4} \mathrm{GeV}$

The study of the purely leptonic decay $B^{+} \rightarrow \tau^{+} \nu$ [1] is of particular interest because it is sensitive to the product of the $B$ meson decay constant $f_{B}$, and the absolute value of Cabibbo-Kobayashi-Maskawa matrix element $V_{u b}$ [2, 3]. In the Standard Model (SM), the decay proceeds via quark annihilation into a $W^{+}$boson, with a branching fraction given by:

$$
\begin{equation*}
\mathcal{B}\left(B^{+} \rightarrow \tau^{+} \nu\right)=\frac{G_{F}^{2} m_{B} m_{\tau}^{2}}{8 \pi}\left[1-\frac{m_{\tau}^{2}}{m_{B}^{2}}\right]^{2} \tau_{B^{+}} f_{B}^{2}\left|V_{u b}\right|^{2} \tag{1}
\end{equation*}
$$

where $G_{F}$ is the Fermi constant, $\tau_{B^{+}}$is the $B^{+}$lifetime, and $m_{B}$ and $m_{\tau}$ are the $B^{+}$meson and $\tau$ lepton masses. Using $\left|V_{u b}\right|=(4.31 \pm 0.30) \times 10^{-3}$ from experimental measurements of semileptonic B decays [4] and $f_{B}=0.216 \pm 0.022 \mathrm{GeV}$ from lattice QCD [5], the SM estimate of the branching fraction for $B^{+} \rightarrow \tau^{+} \nu$ is $(1.5 \pm 0.4) \times 10^{-4}$.

The process $B^{+} \rightarrow \tau^{+} \nu$ is also sensitive to extensions of the SM. For instance, in two-Higgs doublet models [6] and in the MSSM [7, 8] it could be mediated by charged Higgs bosons. The branching fraction measurement can therefore also be used to constrain the parameter space of extensions to the SM.

The $B^{+} \rightarrow \mu^{+} \nu$ and $B^{+} \rightarrow e^{+} \nu$ decays are significantly helicity suppressed with respect to the $B^{+} \rightarrow$ $\tau^{+} \nu$ channel. However, a search for $B^{+} \rightarrow \tau^{+} \nu$ is experimentally more challenging, due to the presence of multiple neutrinos in the final state, which makes the experimental signature less distinctive. In a previously published analysis using a sample of $383 \times 10^{6}$ $\Upsilon(4 S) \rightarrow B \bar{B}$ decays, based on the reconstruction of a semileptonic $B$ decay on the tag side, the $B A B A R$ collaboration set an upper limit $\mathcal{B}\left(B^{+} \rightarrow \tau^{+} \nu\right)<1.7 \times$ $10^{-4}$ at the $90 \%$ confidence level (CL) [9]. The Belle Collaboration has reported evidence from a search for this decay and the branching fraction was measured to be $\mathcal{B}\left(B^{+} \rightarrow \tau^{+} \nu\right)=\left(1.79_{-0.49}^{+0.56}(\text { stat. })_{-0.51}^{+0.46}\right.$ (syst. $\left.)\right) \times 10^{-4}$ [10].

The data used in this analysis were collected with the BABAR detector at the PEP-II storage ring. The sample corresponds to an integrated luminosity of $346 \mathrm{fb}^{-1}$ at the $\Upsilon(4 S)$ resonance (on-resonance) and $36.3 \mathrm{fb}^{-1}$ taken at 40 MeV below the $\Upsilon(4 S)$ resonance (off-resonance). The on-resonance sample contains $383 \times 10^{6} B \bar{B}$ decays. The detector is described in detail elsewhere [11]. Chargedparticle trajectories are measured in the tracking sys-

[^1]tem composed of a five-layer silicon vertex detector and a 40-layer drift chamber (DCH), operating in a 1.5 T solenoidal magnetic field. A Cherenkov detector is used for $\pi-K$ discrimination, a CsI calorimeter for photon detection and electron identification, and the flux return of the solenoid, which consists of layers of steel interspersed with resistive plate chambers or limited streamer tubes, for muon and neutral hadron identification.

In order to estimate signal selection efficiencies and to study physics backgrounds, we use a BABAR Monte Carlo (MC) simulation based on GEANT4 [12]. In MC simulated signal events one $B^{+}$meson decays to $\tau^{+} \nu$ and the other into any final state. The $B \bar{B}$ and continuum MC samples are, respectively, equivalent to approximatively three times and 1.5 times the accumulated data sample. Beam-related background and detector noise are taken from data and overlaid on the simulated events.

We reconstruct an exclusive decay of one of the $B$ mesons in the event $(\operatorname{tag} B)$ and examine the remaining particle(s) for the experimental signature of $B^{+} \rightarrow \tau^{+} \nu$. In order to avoid experimenter bias, the signal region in data is blinded until the final yield extraction is performed.

The $\operatorname{tag} B$ candidate is reconstructed in the set of hadronic $B$ decay modes $B^{-} \rightarrow D^{(*) 0} X^{-}$1], where $X^{-}$denotes a system of charged and neutral hadrons with total charge -1 composed of $n_{1} \pi^{ \pm}, n_{2} K^{ \pm}, n_{3} K_{S}^{0}$, $n_{4} \pi^{0}$, where $n_{1}+n_{2} \leq 5, n_{3} \leq 2$, and $n_{4} \leq 2$. We reconstruct $D^{* 0} \rightarrow D^{0} \pi^{0}, D^{0} \gamma ; D^{0} \rightarrow K^{-} \pi^{+}, K^{-} \pi^{+} \pi^{0}$, $K^{-} \pi^{+} \pi^{-} \pi^{+}, K_{S}^{0} \pi^{+} \pi^{-}$and $K_{S}^{0} \rightarrow \pi^{+} \pi^{-}$. The kinematic consistency of tag $B$ candidates is checked with the beam energy-substituted mass $m_{\mathrm{ES}}=\sqrt{s / 4-\vec{p}_{B}^{2}}$ and the energy difference $\Delta E=E_{B}-\sqrt{s} / 2$. Here $\sqrt{s}$ is the total energy in the $\Upsilon(4 S)$ center-of-mass (CM) frame, and $\vec{p}_{B}$ and $E_{B}$ denote, respectively, the momentum and energy of the tag $B$ candidate in the CM frame. The resolution on $\Delta E$ is measured to be $\sigma_{\Delta E}=10-35 \mathrm{MeV}$, depending on the decay mode; we require $|\Delta E|<3 \sigma_{\Delta E}$. The purity $\mathcal{P}$ of each reconstructed $B$ decay mode is estimated, using on-resonance data, as the ratio of the number of peaking events with $m_{\mathrm{ES}}>5.27 \mathrm{GeV} / c^{2}$ to the total number of events in the same range. If multiple tag $B$ candidates are reconstructed, the one with the highest purity $\mathcal{P}$ is selected. If more than one candidate with the same purity is reconstructed, the one with the lowest value of $|\Delta E|$ is selected. From the dataset obtained as described above, we consider only those events in which the $\operatorname{tag} B$ is reconstructed in the decay modes of highest purity $\mathcal{P}$. The set of decay modes used is defined by the requirement that the purity of the resulting sample is not less than $30 \%$.

The background consists of $e^{+} e^{-} \rightarrow q \bar{q}(q=u, d, s, c)$ events and other $\Upsilon(4 S) \rightarrow B^{0} \bar{B}^{0}$ or $B^{+} B^{-}$decays in which the tag $B$ candidate is mis-reconstructed using particles coming from both $B$ mesons in the event. To reduce
the $e^{+} e^{-} \rightarrow q \bar{q}$ background, we require $\left|\cos \theta_{T B}^{*}\right|<0.9$, where $\theta_{T B}^{*}$ is the angle in the CM frame between the thrust axis [13] of the tag $B$ candidate and the thrust axis of the remaining reconstructed charged and neutral candidates.

In order to determine the number of correctly reconstructed $B^{+}$decays, we classify the background events in four categories: $e^{+} e^{-} \rightarrow c \bar{c} ; e^{+} e^{-} \rightarrow u \bar{u}, d \bar{d}, s \bar{s} ; \Upsilon(4 S)$ $\rightarrow B^{0} \bar{B}^{0}$; and $\Upsilon(4 S) \rightarrow B^{+} B^{-}$. The $m_{\mathrm{ES}}$ shapes of these background distributions are taken from MC simulation. The normalization of the $e^{+} e^{-} \rightarrow c \bar{c}$ and $e^{+} e^{-} \rightarrow$ $u \bar{u}, d \bar{d}, s \bar{s}$ backgrounds is taken from off-resonance data, scaled by the luminosity and corrected for the different selection efficiencies evaluated with the MC. The normalization of the $B^{0} \bar{B}^{0}, B^{+} B^{-}$components are obtained by means of a $\chi^{2}$ fit to the $m_{E S}$ distribution in the data sideband region $\left(5.22 \mathrm{GeV} / c^{2}<m_{\mathrm{ES}}<5.26 \mathrm{GeV} / c^{2}\right)$. The number of background events in the signal region ( $m_{\mathrm{ES}}>$ $5.27 \mathrm{GeV} / c^{2}$ ) is extrapolated from the fit and subtracted from the data. We estimate the total number of tagged $B$ 's in the data to be $N_{B}=(5.92 \pm 0.11$ (stat) $) \times 10^{5}$. Figure 1 shows the tag $B$ candidate $m_{\mathrm{ES}}$ distribution, with the combinatorial background, estimated as the sum of the four components described above, overlaid.


FIG. 1: Distribution of the energy substituted mass, $m_{\mathrm{ES}}$, of the tag $B$ candidates in data. The combinatorial background is overlaid.

After the reconstruction of the tag $B$ meson, a set of selection criteria is applied to the rest of the event (recoil) in order to enhance the sensitivity to $B^{+} \rightarrow \tau^{+} \nu$ decays. We require the presence of only one well-reconstructed charged track (signal track) with charge opposite to that of the tag $B$. The signal track is required to have at least 12 hits in the DCH , momentum transverse to the beam axis, $p_{T}$, greater than $0.1 \mathrm{GeV} / c$, and the point of closest approach to the interaction point less than 10 cm along the beam axis and less than 1.5 cm transverse to it.

The $\tau$ lepton is identified in four decay modes constituting approximately $71 \%$ of the total $\tau$ decay width: $\tau^{+} \rightarrow e^{+} \nu \bar{\nu}, \tau^{+} \rightarrow \mu^{+} \nu \bar{\nu}, \tau^{+} \rightarrow \pi^{+} \bar{\nu}$, and $\tau^{+} \rightarrow \pi^{+} \pi^{0} \bar{\nu}$. Particle identification criteria on the signal track are used to separate the four categories. The $\tau^{+} \rightarrow \pi^{+} \pi^{0} \bar{\nu}$ sample
is obtained by associating the signal track, identified as pion, with a $\pi^{0}$ reconstructed from a pair of neutral clusters with invariant mass between 0.115 and $0.155 \mathrm{GeV} / c^{2}$ and total energy greater than 250 MeV . In case of multiple $\pi^{+} \pi^{0}$ candidates, the one with largest center-of-mass momentum $p_{\pi^{+} \pi^{0}}^{*}$ is chosen.

We place a mode-dependent cut on $\left|\cos \theta_{T B}^{*}\right|$ to reduce the background due to continuum events and incorrectly reconstructed tag $B$ candidates (combinatorial). The remaining sources of background consists of $B^{+} B^{-}$events in which the tag $B$ meson was correctly reconstructed and the recoil contains one track and additional particles that are not reconstructed by the tracking detectors and calorimeter. MC simulation shows that most of this background is from semileptonic $B$ decays.

We define the discriminating variable $E_{\text {extra }}$ as the sum of the energies of the neutral clusters not associated with the tag $B$ or with the signal $\pi^{0}$ from the $\tau^{+} \rightarrow \pi^{+} \pi^{0} \bar{\nu}$ mode, and passing a minimum energy requirement. The required energy depends on the selected signal mode and on the calorimeter region involved and varies from 50 to 70 MeV . Signal events tend to peak at low $E_{\text {extra }}$ values, whereas background events, which contain additional sources of neutral clusters, are distributed toward higher $E_{\text {extra }}$ values.

Other variables used to discriminate between signal and background are the CM momentum of the signal candidates, the multiplicities of low $p_{T}$ charged tracks and of $\pi^{0}$ candidates in the recoil, and the direction of the missing momentum four-vector in the CM frame. For the $\tau^{+} \rightarrow \pi^{+} \pi^{0} \bar{\nu}$ mode, we exploit the presence of the $\pi^{0}$ in the final state and the dominance of the decay through the $\rho^{+}$resonance by means of the combined quantity $x_{\rho}=\left[\left(m_{\pi^{+} \pi^{0}}-m_{\rho}\right) /\left(\Gamma_{\rho}\right)\right]^{2}+\left[\left(m_{\gamma \gamma}-m_{\pi^{0}}\right) /\left(\sigma_{\pi^{0}}\right)\right]^{2}$, where $m_{\pi^{+} \pi^{0}}$ is the reconstructed invariant mass of the $\pi^{+} \pi^{0}$ candidate, $m_{\gamma \gamma}$ is the reconstructed invariant mass of the $\pi^{0}$ candidate, $m_{\rho}$ and $\Gamma_{\rho}$ are the nominal values [4] for the $\rho$ mass and width, $m_{\pi^{0}}$ is the nominal $\pi^{0}$ mass and $\sigma_{\pi^{0}}=8 \mathrm{MeV} / c^{2}$ is the experimental resolution on the $\pi^{0}$ mass determined from data.

We optimize the selection by maximizing $s / \sqrt{s+b}$ using the $B^{+} B^{-}$MC and signal MC, where $b$ is the expected background from $B^{+} B^{-}$events and $s$ is the expected number of signal events in the hypothesis of a branching fraction of $1 \times 10^{-4}$. The optimization is performed separately for each $\tau$ decay mode and with all the cuts applied simultaneously in order to take into account any correlations among the discriminating variables. The optimized signal selection cuts are reported in Table $\mathbb{I}$

We compute the signal selection efficiency as the ratio of the number of signal MC events passing the selection criteria to the number of signal events that have a correctly reconstructed tag $B$ candidate in the signal region $m_{\mathrm{ES}}>5.27 \mathrm{GeV} / c^{2}$. We evaluate the efficiencies on a signal MC sample which is distinct from the sample used in the optimization procedure. A small cross-feed in some modes is estimated from MC and is taken into account in the computation of the total efficiency.

TABLE I: Optimized selection criteria for each $\tau$ decay mode.

| Variable | $e^{+}$ | $\mu^{+}$ | $\pi^{+}$ | $\pi^{+} \pi^{0}$ |
| :--- | ---: | ---: | ---: | ---: |
| $E_{\text {extra }}(\mathrm{GeV})$ | $<0.160$ | $<0.100$ | $<0.230$ | $<0.290$ |
| $\pi^{0}$ multiplicity | 0 | 0 | $\leq 2$ | - |
| Track multiplicity | 1 | 1 | $\leq 2$ | 1 |
| $\left\|\cos \theta_{T B}^{*}\right\|$ | $\leq 0.9$ | $\leq 0.9$ | $\leq 0.7$ | $\leq 0.7$ |
| $p_{\text {trk }}^{*}(\mathrm{GeV} / c)$ | $<1.25$ | $<1.85$ | $>1.5$ | - |
| $\cos \theta_{\text {miss }}^{*}$ | $<0.9$ | - | $<0.5$ | $<0.55$ |
| $p_{\pi^{+} \pi^{0}}^{*}(\mathrm{GeV} / c)$ | - | - | - | $>1.5$ |
| $x_{\rho}$ | - | - | - | $<2.0$ |
| $E_{\pi^{0}}(\mathrm{GeV})$ | - | - | $->0.250$ |  |

The total efficiency for each selection is given by:

$$
\begin{equation*}
\varepsilon_{i}=\sum_{j=1}^{n_{\mathrm{dec}}} \varepsilon_{i}^{j} f_{j} \tag{2}
\end{equation*}
$$

where $\varepsilon_{i}^{j}$ is the efficiency of the selection $i$ for the $\tau$ decay mode $j, n_{\text {dec }}=7$ is the number of $\tau$ decay modes that can contribute to the reconstructed modes and $f_{j}$ are the fractions of the $\tau$ decay mode as estimated from the signal MC sample with a reconstructed tag $B$. Table $\Pi$ shows the estimated efficiencies.

TABLE II: Efficiency (in percent) of the most relevant $\tau$ decay modes (rows) to be selected in one of the four modes considered in this analysis (column). The All decay row shows the selection efficiency of each reconstruction mode, adding the contribution from the previous rows, weighted by the decay abundance at the tag selection level $f_{j}$. The last row shows the total signal selection efficiency. The uncertainties are statistical only.

| Mode | $e^{+}$ | $\mu^{+}$ | $\pi^{+}$ | $\pi^{+} \pi^{0}$ |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $e^{+}$ | $\mathbf{1 9 . 3} \pm \mathbf{1 . 1}$ | 0 | $0.4 \pm 0.2$ | 0 |  |
| $\mu^{+}$ | 0 | $\mathbf{1 0 . 8} \pm \mathbf{0 . 9}$ | $1.3 \pm 0.3$ | 0 |  |
| $\pi^{+}$ | 0 | $0.1 \pm 0.1$ | $\mathbf{1 9 . 7} \pm \mathbf{1 . 3}$ | $0.5 \pm 0.2$ |  |
| $\pi^{+} \pi^{0}$ | 0 | 0 | $1.5 \pm 0.2$ | $\mathbf{7 . 0} \pm \mathbf{0 . 5}$ |  |
| $\pi^{+} \pi^{+} \pi^{-}$ | 0 | 0 | 0 | 0 |  |
| $\pi^{+} \pi^{0} \pi^{0}$ | 0 | 0 | $0.2 \pm 0.1$ | $1.8 \pm 0.4$ |  |
| Other | 0 | 0 | $0.3 \pm 0.2$ | $0.1 \pm 0.1$ |  |
| All dec. $\epsilon_{i}:$ | $\mathbf{3 . 1} \pm \mathbf{0 . 2}$ | $\mathbf{1 . 7} \pm \mathbf{0 . 1}$ | $\mathbf{2 . 9} \pm \mathbf{0 . 2}$ | $\mathbf{2 . 2} \pm \mathbf{0 . 2}$ |  |
| Total: |  |  |  |  |  |

To determine the expected number of background events in the data, we use the final selected data samples with $E_{\text {extra }}$ between 0 and 2.4 GeV . We first perform an extended unbinned maximum likelihood fit to the $m_{\mathrm{ES}}$ distribution in the $E_{\text {extra }}$ sideband region $0.4 \mathrm{GeV}<E_{\text {extra }}<2.4 \mathrm{GeV}$ of the final sample. For the peaking component of the background we use a probability density function (PDF) which is a Gaussian function joined to an exponential tail (Crystal Ball function) 14].

TABLE III: Observed number of on-resonance data events in the signal region compared with the number of expected background events.

| $\tau$ decay mode | Expected background | Observed |
| :--- | :---: | :---: |
| $\tau^{+} \rightarrow e^{+} \nu \bar{\nu}$ | $1.47 \pm 1.37$ | 4 |
| $\tau^{+} \rightarrow \mu^{+} \nu \bar{\nu}$ | $1.78 \pm 0.97$ | 5 |
| $\tau^{+} \rightarrow \pi^{+} \bar{\nu}$ | $6.79 \pm 2.11$ | 10 |
| $\tau^{+} \rightarrow \pi^{+} \pi^{0} \bar{\nu}$ | $4.23 \pm 1.39$ | 5 |
| All modes | $14.27 \pm 3.03$ | 24 |

As a PDF for the non-peaking component, we use a phase space motivated threshold function (ARGUS function) [15]. From this fit, we determine a peaking yield $N_{p k}^{\text {side,data }}$ and signal shape parameters, to be used in later fits. We apply the same procedure to $B^{+} B^{-}$MC events which pass the final selection and determine the peaking yield $N_{p k}^{\text {side,MC }}$. To determine the MC peaking yield in the $E_{\text {extra }}$ signal region $N_{p k}^{\mathrm{sig}, \mathrm{MC}}$, we fit $m_{\mathrm{ES}}$ in the $E_{\text {extra }}$ signal region of the $B^{+} B^{-}$MC sample with the Crystal Ball parameters fixed to the values determined in the $E_{\text {extra }}$ sideband fits described above. Analogously, we fit the $m_{\mathrm{ES}}$ distribution of data in the $E_{\text {extra }}$ signal region to extract the combinatorial background $n_{\text {comb }}$, evaluated as the integral of the ARGUS shaped component in the $m_{\mathrm{ES}}>5.27 \mathrm{GeV} / c^{2}$ region. We estimate the total expected background in the signal region as:

$$
\begin{equation*}
b=\frac{N_{p k}^{\mathrm{sig}, \mathrm{MC}}}{N_{p k}^{\text {side,MC }}} \times N_{p k}^{\text {side,data }}+n_{\mathrm{comb}} \tag{3}
\end{equation*}
$$

After finalizing the signal selection criteria, we measure the yield of events in each decay mode in on-resonance data. Table IIIreports the number of observed events together with the expected number of background events, for each $\tau$ decay mode. Figure 2 shows the $E_{\text {extra }}$ distribution for data and expected background at the end of the selection. The signal MC, normalized to a branching fraction of $3 \times 10^{-3}$ for illustrative purposes, is overlaid for comparison. The $E_{\text {extra }}$ distribution is also plotted separately for each $\tau$ decay mode.

We combine the results on the observed number of events $n_{i}$ and on the expected background $b_{i}$ from each of the four signal decay modes $\left(n_{c h}\right)$ using the estimator $Q=\mathcal{L}(s+b) / \mathcal{L}(b)$, where $\mathcal{L}(s+b)$ and $\mathcal{L}(b)$ are the likelihood functions for signal plus background and background-only hypotheses, respectively:

$$
\begin{equation*}
\mathcal{L}(s+b) \equiv \prod_{i=1}^{n_{c h}} \frac{e^{-\left(s_{i}+b_{i}\right)}\left(s_{i}+b_{i}\right)^{n_{i}}}{n_{i}!}, \mathcal{L}(b) \equiv \prod_{i=1}^{n_{c h}} \frac{e^{-b_{i}} b_{i}^{n_{i}}}{n_{i}!} \tag{4}
\end{equation*}
$$

The estimated number of signal candidates $s_{i}$ in data, for each decay mode, is related to the $B^{+} \rightarrow \tau^{+} \nu$ branching


FIG. 2: $E_{\text {extra }}$ distribution after all selection criteria have been applied. The upper plot shows the distribution of all the modes combined while lower plots show the (a) $\tau^{+} \rightarrow e^{+} \nu \bar{\nu}$, (b) $\tau^{+} \rightarrow \mu^{+} \nu \bar{\nu}$, (c) $\tau^{+} \rightarrow \pi^{+} \bar{\nu}$, and (d) $\tau^{+} \rightarrow \pi^{+} \pi^{0} \bar{\nu}$ modes separately. The on-resonance data (black dots) distribution is compared with the total background prediction (continuous histogram). The hatched histrogram represents the combinatorial background component. $B^{+} \rightarrow \tau^{+} \nu$ signal MC (dashed histogram), normalized to a branching fraction of $3 \times 10^{-3}$ for illustrative purposes, is shown for comparison.
fraction by:

$$
\begin{equation*}
s_{i}=\frac{\varepsilon_{\mathrm{sig}}^{\mathrm{tag}}}{\varepsilon_{B}^{\mathrm{tag}}} N_{B^{+}}^{\mathrm{tag}} \varepsilon_{i} \mathcal{B}\left(B^{+} \rightarrow \tau^{+} \nu\right) \tag{5}
\end{equation*}
$$

where $N_{B^{+}}^{\mathrm{tag}}$ is the number of tag $B^{+}$mesons correctly reconstructed, $\varepsilon_{B}^{\mathrm{tag}}$ and $\varepsilon_{\mathrm{sig}}^{\mathrm{tag}}$ are the tag $B$ efficiencies in generic $B \bar{B}$ and signal events respectively, and $\varepsilon_{i}$ are the signal efficiencies defined in equation 2. We fix the ratio $\varepsilon_{\text {sig }}^{\text {tag }} / \varepsilon_{B}^{\text {tag }}=0.939 \pm 0.007$ (stat.) to the value obtained from MC simulation.

We estimate the branching fraction (including statistical uncertainty and uncertainty from the background) by scanning over signal branching fraction hypotheses and computing the value of $\mathcal{L}(s+b) / \mathcal{L}(b)$ for each hypothesis. The branching fraction is the hypothesis which minimizes the likelihood ratio $-2 \ln \mathrm{Q}=-2 \ln (\mathcal{L}(s+b) / \mathcal{L}(b))$, and we determine the statistical uncertainty by finding the points on the likelihood scan that occur at one unit above the minimum.

The dominant uncertainty on the background predic-
tions $b_{i}$ is due to the finite $B^{+} B^{-}$MC statistics. We also check possible systematic effects in the estimation of combinatorial background by means of a sample of events with looser selection requirements; we find it to be negligible with respect to the statistical uncertainty. The background uncertainty is incorporated in the likelihood definition used to extract the branching fraction, by convolving it with a Gaussian function with standard deviation equal to the error on $b_{i}$ [16].

The other sources of systematic uncertainty in the determination of the $B^{+} \rightarrow \tau^{+} \nu$ branching fraction come from the estimation of the tag yield and efficiency and the reconstruction efficiency of the signal modes. We estimate the systematic uncertainty on the tag $B$ yield and reconstruction efficiency by varying the MC $B^{+} B^{-}$nonpeaking component of the $m_{\mathrm{ES}}$ shape, assigning a systematic uncertainty of $3 \%$ on the branching fraction. The systematic uncertainties due to mismodeling of charged particle tracking efficiency, $E_{\text {extra }}$ shape, particle identification efficiency, $\pi^{0}$ reconstruction and signal MC statistics depend on the $\tau$ decay mode. The uncertainty on the branching fraction is evaluated for each mode separately. We obtain the total contributions due to tracking and $E_{\text {extra }}$ systematics by adding linearly the contributions of each decay channel. The total contributions due to MC statistics and particle identification are obtained by adding systematics uncertainties of each reconstruction mode in quadrature.

We check the low $p_{T}$ charged track multiplicity distribution agreement between data and MC with a sample enriched in background by loosening the selection criteria. The disagreement, which is mode dependent, is quantified by comparing the MC PDF with the data PDF. We correct the MC to reproduce the distribution in data and apply the correction to the signal MC distribution. We take $100 \%$ of the correction as a systematic uncertainty, resulting in a total systematic uncertainty of $5.8 \%$ on the branching fraction.

The systematic uncertainty due to the $E_{\text {extra }}$ mismodeling is determined by means of a data sample containing events with two non-overlapping tag $B$ candidates. The sample is selected by reconstructing a second $B$ meson in a hadronic decay mode $B^{-} \rightarrow D^{(*) 0} X^{-}$on the recoil of the $\operatorname{tag} B$. In addition to the requirements on the $\operatorname{tag} B$ described above, we consider only second $B$ candidates satisfying $|\Delta E|<50 \mathrm{MeV}$ and $m_{\mathrm{ES}}>5.27 \mathrm{GeV} / c^{2}$ having opposite charge to that of the tag $B$. If multiple candidates are reconstructed, the one with the highest purity $\mathcal{P}$ is selected. We compare the distribution of the total energy of the unassigned neutral clusters $E_{\text {extra }}$ in data and in MC. We compute the ratio of the number of events in the signal region of each $\tau$ mode to the total number of events in the sample. For each $\tau$ mode, we evaluate the systematic uncertainty, comparing the ratio estimated from MC to the ratio estimated from data. This procedure results in a $8.8 \%$ systematic uncertainty on the branching fraction. Table IV shows the contributions in percent to the systematic uncertainties on the

TABLE IV: Contributions (in percent) to the systematic uncertainty on the branching fraction due to signal selection efficiency for different selection modes.

| Source of systematics | $e^{+}$ | $\mu^{+}$ | $\pi^{+}$ | $\pi^{+}$ | $\pi^{0}$ | Total |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| MC statistics | 3.1 | 0.6 | 1.5 | 2.6 | 4.3 |  |  |
| Particle Identification | 1.5 | 1.3 | 0.2 | 0.2 | 2.0 |  |  |
| $\pi^{0}$ | - | - | - | 1.4 | 1.4 |  |  |
| Tracking | 3.7 | 0.4 | 0.1 | 1.6 | 5.8 |  |  |
| $E_{\text {extra }}$ | 4.7 | 0.6 | 0.9 | 2.6 | 8.8 |  |  |
| Signal $B$ |  |  |  |  |  | 11.6 |  |
| Tag $B$ |  |  |  |  |  |  | 3 |
| Total |  |  |  |  | 12 |  |  |

branching fraction.
In summary, we measure the branching fraction

$$
\begin{equation*}
\mathcal{B}\left(B^{+} \rightarrow \tau^{+} \nu\right)=\left(1.8_{-0.8}^{+0.9} \pm 0.4 \pm 0.2\right) \times 10^{-4} \tag{6}
\end{equation*}
$$

where the first error is statistical, the second is due to the background uncertainty, and the third is due to other systematic sources. Taking into account the uncertainty on the expected background, as described above, we obtain a significance of $2.2 \sigma$.

Using Eq. [1 we calculate the product of
the $B$ meson decay constant $f_{B}$ and $\left|V_{u b}\right|$ to be $f_{B} \cdot\left|V_{u b}\right|=\left(10.1_{-2.5}^{+2.3} \text { (stat. }\right)_{-1.5}^{+1.2}($ syst. $\left.)\right) \times 10^{-4} \mathrm{GeV}$. We also measure the $90 \%$ C.L. upper limit using the $C L_{s}$ method [17] to be $\mathcal{B}\left(B^{+} \rightarrow \tau^{+} \nu\right)<3.4 \times 10^{-4}$.

The combination of this measurement with the $B A B A R$ result obtained using semileptonic tags, based on a statistically independent data sample, and reported in [9], yields:
$\mathcal{B}\left(B^{+} \rightarrow \tau^{+} \nu\right)=\left(1.2 \pm 0.4_{\text {stat. }} \pm 0.3_{\text {bkg. }} \pm 0.2_{\text {syst. }}\right) \times 10^{-4}$.
The significance of the combined result is $2.6 \sigma$ including the uncertainty on the expected background (3.2 $\sigma$ if this uncertainty is not included).

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