## A Search for the Rare Leptonic Decay $B^{-} \rightarrow \tau^{-} \overline{\boldsymbol{\nu}_{\boldsymbol{\tau}}}$

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#### Abstract

We present a search for the decay $B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}$ in a sample of $88.9 \times 10^{6} B \bar{B}$ pairs recorded with the BABAR detector at the SLAC $B$-Factory. One of the two $B$ mesons from the $\Upsilon(4 S)$ is reconstructed in a hadronic or a semileptonic final state and the decay products of the other $B$ in the event are analyzed for consistency with a $B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}$ decay. We find no evidence of a signal and set an upper limit on the branching fraction of $\mathcal{B}\left(B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}\right)<4.2 \times 10^{-4}$ at the $90 \%$ confidence level.


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In the Standard Model (SM) the leptonic decay $B^{-} \rightarrow$ $\tau^{-} \overline{\nu_{\tau}}$ 1] proceeds via the annihilation of the $b$ and $\bar{u}$ quarks into a virtual $W$ boson. Its amplitude is thus proportional to the product of the Cabibbo-KobayashiMaskawa (CKM) matrix 2] element $\left|V_{u b}\right|$ and the $B$ meson decay constant $f_{B}$. The SM branching fraction is given by:

$$
\begin{align*}
\mathcal{B}\left(B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}\right) & =\frac{G_{F}^{2} m_{B}}{8 \pi} m_{\tau}^{2}\left(1-\frac{m_{\tau}^{2}}{m_{B}^{2}}\right)^{2} f_{B}^{2}\left|V_{u b}\right|^{2} \tau_{B} \\
& =(9.3 \pm 3.9) \times 10^{-5} \tag{1}
\end{align*}
$$

where $G_{F}$ is the Fermi coupling constant, $m_{\tau}$ and $m_{B}$ are the $\tau$ lepton and $B^{-}$meson masses, and $\tau_{B}$ is the $B^{-}$mean lifetime. We have used $\tau_{B}=(1.671 \pm 0.018) \mathrm{ps}$, $\left|V_{u b}\right|=(3.67 \pm 0.47) \times 10^{-4}$, and $f_{B}=(0.196 \pm 0.032) \mathrm{GeV}$ (obtained from lattice QCD calculations) [3]. The branching fractions for $e^{-} \bar{\nu}_{e}$ and $\mu^{-} \bar{\nu}_{\mu}$ are helicity suppressed by factors of $\sim 10^{-8}$ and $\sim 10^{-3}$, respectively. Physics beyond the SM, such as supersymmetry or twoHiggs doublet models, could enhance $\mathcal{B}\left(B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}\right)$ by up to a factor of five through the introduction of a charged Higgs boson (4].

A search for this decay is experimentally challenging due to the presence of at least two undetectable neutrinos in the final state. No observation has been reported yet and the most stringent published limit on the decay is $\mathcal{B}\left(B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}\right)<5.7 \times 10^{-4}$ at the $90 \%$ confidence level [5].

The data used in this analysis were recorded with the $B A B A R$ detector at the PEP-II asymmetric $e^{+} e^{-}$storage ring. The sample consists of $88.9 \pm 1.0$ million $B \bar{B}$ pairs $\left(81.9 \mathrm{fb}^{-1}\right)$ collected at the $\Upsilon(4 S)$ resonance ("onresonance") and $9.6 \mathrm{fb}^{-1}$ collected about 40 MeV below the $B \bar{B}$ threshold ("off-resonance").

The $B A B A R$ detector is described in detail elsewhere [6]. Detection of charged particles and measurement of their momenta are performed by a five-layer double-sided silicon vertex tracker and a 40-layer drift chamber, which operate in a $1.5-\mathrm{T}$ solenoidal magnetic field. A detector of internally reflected Cherenkov light is used to identify charged kaons and pions. Photons and electrons are detected in an electromagnetic calorimeter consisting of an array of $\mathrm{CsI}(\mathrm{Tl})$ crystals. Muons and neutral hadrons are identified in the flux return, which is instrumented with multiple layers of resistive plate chambers. A GEANT4based [7] simulation of the $B A B A R$ detector, including
machine backgrounds, is used to study signal event selection and background rejection.

We first select a sample of events with one $B$-meson (the $\operatorname{tag} B$ ) reconstructed in a hadronic or a semileptonic final state. The reconstruction constrains the kinematics and reduces the combinatorics in each event. This is critical since at least two neutrinos result from the $B^{-} \rightarrow$ $\tau^{-} \overline{\nu_{\tau}}$ decay. All the neutral and charged particles not used for the tag $B$ are assumed to come from the $B$ meson recoiling against it. We use two methods to search this recoil system for evidence of a $B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}$ signal.

In our first method, we reconstruct the tag $B$ semileptonically. The semileptonic $B$-meson, $B_{\mathrm{sl}}$, is reconstructed as $B^{+} \rightarrow \bar{D}^{0} \ell^{+} \nu_{\ell} X$, where $\ell=e, \mu$ and $X$ can be a $\gamma, \pi^{0}$, or nothing. We select semileptonic $B$-decay events with several missing particles (such as neutrinos) by requiring at least one lepton with center-of-mass (CM) momentum $\left(\left|\vec{p}_{\ell}^{*}\right|\right)$ above $1.0 \mathrm{GeV} / c$, zero event charge, a ratio of the Fox-Wolfram moments [8] $H_{2} / H_{0}<0.9$, and missing mass greater than $1.0 \mathrm{GeV} / c^{2}$. Here, the missing mass is determined by subtracting the total energy and momentum of all reconstructed tracks and neutrals from the four-momentum of the $\Upsilon(4 S)$ system. We reconstruct $\bar{D}^{0}$ mesons in the modes $\bar{D}^{0} \rightarrow$ $K^{+} \pi^{-}, K^{+} \pi^{-} \pi^{-} \pi^{+}, K^{+} \pi^{-} \pi^{0}$, and $K_{S}^{0} \pi^{+} \pi^{-}$and require their reconstructed masses to be within three standard deviations of the observed mean. The $\bar{D}^{0}$ mesons are then paired with leptons with $\left|\vec{p}_{\ell}^{*}\right|>1.0 \mathrm{GeV} / c$ to form $D \ell$ candidates. If the $\bar{D}^{0}$ decay contains a charged kaon, the lepton must have the same charge as the kaon. The $\bar{D}^{0}$ and lepton are required to originate from a common vertex, but we do not mass-constrain the vertex fit. We assume that the only missing particle is a neutrino and calculate the cosine of the angle between the momentum vectors of the $D \ell$ candidate and the $B$-meson,

$$
\begin{equation*}
\cos \theta_{B, D \ell} \equiv \frac{2 E_{\text {beam }}^{*} E_{D \ell}^{*}-m_{B}^{2}-m_{D \ell}^{2}}{2 \sqrt{E_{\text {beam }}^{* 2}-m_{B}^{2}}\left|\vec{p}_{D \ell}^{*}\right|} \tag{2}
\end{equation*}
$$

The CM energy and momentum of the $D \ell$ candidate are $E_{D \ell}^{*}$ and $\vec{p}_{D \ell}^{*}$, respectively. The $B$-meson energy is taken to be the beam energy, $E_{\text {beam. }}^{*}$. Calculated values of $\cos \theta_{B, D \ell}$ may lie outside the physical range for events where the $D \ell$ candidate did not arise as presumed, or due to detector energy and momentum resolution. We place an asymmetric restriction on this variable, $-2.5<\cos \theta_{B, D \ell}<1.1$, to admit $\bar{D}^{* 0}$ states where additional decay products are present. If there is more than


FIG. 1: The distribution of $E_{\text {extra }}$ after applying all selection criteria. The fit to the data and its components are also shown. The background is normalized to the data luminosity and the signal simulation is normalized arbitrarily.
one acceptable $D \ell$ candidate, we choose the one whose $\bar{D}^{0}$ mass is closest to the mean of the fitted distribution.

After identifying the $B_{\mathrm{sl}}$, the remaining particles are required to be consistent with $B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}$ where $\tau^{-} \rightarrow$ $e^{-} \bar{\nu}_{e} \nu_{\tau}$ or $\mu^{-} \bar{\nu}_{\mu} \nu_{\tau}$. Exactly one track with a small impact parameter relative to the primary vertex must remain. The track must have $p^{*}<1.2 \mathrm{GeV} / c$, and must be identified as either an electron or muon. We reject $e^{+} e^{-} \rightarrow \tau^{+} \tau^{-}$events by restricting the angle of the track with respect to the event thrust axis $\left(\left|\cos \theta_{\vec{p}, \vec{T}}\right|<0.9\right)$ and the minimum invariant mass constructable from any triplet of tracks in the event $\left(M_{3}^{\min }>1.5 \mathrm{GeV} / c^{2}\right)$. In general, continuum events tend to peak sharply at $\left|\cos \theta_{\vec{p}, \vec{T}}\right|=1$ and $\tau^{+} \tau^{-}$events in particular tend to peak at values of $M_{3}^{\min }$ below the $\tau$ mass.

The signal yield in the data is determined using the distribution of the total energy deposited in calorimeter clusters (with a minimum energy of 0.020 GeV ) by neutral particles not associated with the $\bar{D}^{0}$ decay in the semileptonic $B_{\text {sl }}$ candidate, $E_{\text {extra }}$ (Fig. (1). This variable peaks near zero for signal while for background it rises with increasing $E_{\text {extra }}$. For $E_{\text {extra }}<1.0 \mathrm{GeV}$, we find from Monte Carlo simulations a signal efficiency of $(4.77 \pm 0.35) \times 10^{-4}$ and a background estimate of $124 \pm 7$ events.

The signal efficiency quoted above is determined using a detailed signal simulation. We study the differences between simulation and data in the semileptonic $B$ reconstruction, neutral-energy reconstruction, and lepton identification to derive an efficiency correction. The most significant effect comes from the $B_{\text {sl }}$ reconstruction efficiency, and is determined using a sample of events in data and Monte Carlo simulations where both $B$ mesons are reconstructed as $B \rightarrow D \ell \nu X$. The total
efficiency correction from all sources is determined to be $0.878 \pm 0.076$, and the corrected signal efficiency is $\left(4.19 \pm 0.31_{\text {stat }} \pm 0.36_{\text {syst }}\right) \times 10^{-4}$.

Probability density functions (p.d.f.'s) are constructed from the $E_{\text {extra }}$ distributions in signal $\left(F\left(E_{\text {extra }}\right)_{s}\right)$ and background $\left(F\left(E_{\text {extra }}\right)_{b}\right)$ simulations. The $E_{\text {extra }}$ distribution for signal events is modeled as the sum of an exponential and two Gaussian distributions. The double-Gaussian models signal events where the $X$ in $B^{+} \rightarrow \bar{D}^{0} \ell^{+} \nu_{\ell} X$ is a $\pi^{0}$ or photon with a characteristic energy around 0.15 GeV . The exponential models signal events where such neutral particles are absent. To model background, as determined from Monte Carlo, we use a third-order polynomial. The p.d.f.'s are combined into an extended maximum likelihood function,

$$
\begin{equation*}
\mathcal{L}(s+b) \equiv \frac{e^{-\mu_{s}-\mu_{b}}}{n!} \prod_{i=1}^{n}\left[\mu_{s} F\left(E_{i}\right)_{s}+\mu_{b} F\left(E_{i}\right)_{b}\right] \tag{3}
\end{equation*}
$$

where $E_{i}$ is the $E_{\text {extra }}$ in the $i$ th event, $n$ is the total number of events in the data, and $\mu_{s}$ and $\mu_{b}$ are the signal and background yields to be fitted in the data. Studies of the choice of p.d.f. parameterization and of variations in shape suggest that the chosen p.d.f.'s yield a consistently conservative limit for the upper bound of the branching fraction. We fix the p.d.f. shape parameters and fit the data (Fig. [1). The fit yields $14.8 \pm 6.3$ signal events and $115.2 \pm 11.8$ background events. This signal yield has a statistical significance of $2.3 \sigma$.

We set a limit on the branching fraction at the $90 \%$ confidence level (C.L.) using the "CLs method" described in Refs. [9] 10]. We define our statistical estimator, $Q$, to be the fitted signal yield and compare the value of $Q$ in data to its value in a large number of experiments generated by sampling the likelihood function over a range of signal hypotheses. The uncertainty in the signal efficiency estimate is included by assuming a Gaussian uncertainty in the signal hypothesis. Using our fitted signal yield, efficiency, and the total number of $B$ mesons in the data sample we determine that $\mathcal{B}\left(B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}\right)<$ $6.7 \times 10^{-4}(90 \%$ C.L. $)$.

In our second method, we reconstruct the tag $B$ candidate, $B_{\text {had }}$, decaying into a set of purely hadronic final states, $B^{+} \rightarrow \bar{D}^{(*) 0} X^{+}$. The $\bar{D}^{* 0}$ is reconstructed in the mode $\bar{D}^{0} \pi^{0}$, and $X^{+}$is a system of hadrons composed of $n_{1} \pi^{ \pm}+n_{2} K^{ \pm}+n_{3} \pi^{0}+n_{4} K_{S}^{0}$ where $n_{1}=1, \ldots 5$; $n_{2}=0,1,2 ; n_{3}=0,1,2$; and $n_{4}=0,1$. Rejection of background processes is based on two kinematic quantities: $\Delta E$, the difference between the $B_{\text {had }}$ and beam energies, and the beam-energy-substituted mass $m_{\mathrm{ES}}$,

$$
\begin{equation*}
m_{\mathrm{ES}} \equiv \sqrt{\left[\left(s / 2+\vec{p} \cdot \vec{p}_{B}\right)^{2} / E^{2}\right]-\left|\vec{p}_{B}\right|^{2}} \tag{4}
\end{equation*}
$$

where $\sqrt{s}$ is the total energy of the $e^{+} e^{-}$system in the CM frame, and $(E, \vec{p})$ and $\left(E_{B}, \vec{p}_{B}\right)$ are the fourmomenta of the $e^{+} e^{-}$system and the $B_{\text {had }}$, respectively, both in the laboratory frame.


FIG. 2: Distribution of $m_{\mathrm{ES}}$ for the $B_{\mathrm{had}}$ candidates in data. The events lie in the region $-0.1<\Delta E<0.08 \mathrm{GeV}$. The solid curve shows the result of the fit with the sum of a Crystal Ball function (dashed curve) and an ARGUS function (dotted curve).

For each mode the $m_{\mathrm{ES}}$ distribution of the reconstructed candidates with $-0.1<\Delta E<0.08 \mathrm{GeV}$ and $m_{\mathrm{ES}}>5.21 \mathrm{GeV} / c^{2}$ is fitted using the sum of a "Crystal Ball function" [11] to model the signal component peaking at $m_{B}$ and an "ARGUS function" 12] to model the continuum and combinatorial $B$ background. Figure 2 shows the fit to the $m_{\mathrm{ES}}$ distribution for the $B_{\text {had }}$ candidates in data. We define the signal region as $-0.09<\Delta E<0.06 \mathrm{GeV}$ and $m_{\mathrm{ES}}>5.27 \mathrm{GeV} / c^{2}$. We define a sideband region, $5.21<m_{\mathrm{ES}}<5.26 \mathrm{GeV} / c^{2}$, to provide a control sample for studying continuum and combinatorial $B$ background. The yield in the signal region, as determined from the fit, is $N_{B_{\text {had }}}=(167.8 \pm$ $\left.1.2_{\text {stat }} \pm 3.0_{\text {syst }}\right) \times 10^{3}$. The error is dominated by systematic uncertainty in the functional form of the peak at $m_{B}$.

We identify the $\tau$ lepton using the following decay channels: $\tau^{-} \rightarrow e^{-} \bar{\nu}_{e} \nu_{\tau}, \mu^{-} \nu_{\tau} \bar{\nu}_{\mu}, \pi^{-} \nu_{\tau}, \pi^{-} \pi^{0} \nu_{\tau}$, and $\pi^{-} \pi^{+} \pi^{-} \nu_{\tau}$. We require the charged particles to be identified as leptons or pions, as appropriate. Modespecific constraints are placed on the particles recoiling against the $B_{\text {had }}$. For the lepton and single-pion modes we reject events with $\pi^{0}$ or $K_{S}^{0}$ mesons in the recoil. The event is required to have zero charge and, in the recoil, at most one photon candidate not associated with a $\pi^{0}$. Events with such a photon candidate are accepted only if $50<E_{\gamma}<100 \mathrm{MeV}(50<$ $E_{\gamma}<110 \mathrm{MeV}$ for the $\tau^{-} \rightarrow e^{-} \bar{\nu}_{e} \nu_{\tau}, \mu^{-} \nu_{\tau} \bar{\nu}_{\mu}$, and $\pi \nu_{\tau}$ modes) in the laboratory frame. Further requirements are made on the total missing momentum of the event, $p_{\text {miss }}>1.2 \mathrm{GeV} / c\left(>1.4 \mathrm{GeV} / c\right.$ for $\left.\tau^{-} \rightarrow \pi^{-} \pi^{0} \nu_{\tau}\right)$, the total momentum of the track(s) in the parent- $B$ rest frame $\left(p_{\pi^{-}}>1.2 \mathrm{GeV} / c\right.$ for $\tau^{-} \rightarrow \pi^{-} \nu_{\tau}, p_{\pi^{-} \pi^{+} \pi^{-}}>$ $1.6 \mathrm{GeV} / c$ for $\tau^{-} \rightarrow \pi^{-} \pi^{+} \pi^{-} \nu_{\tau}$ ), and the invariant mass
of two or three pions $\left(0.60<m_{\pi \pi}<0.95 \mathrm{GeV} / c^{2}\right.$ and $1.10<m_{\pi \pi \pi}<1.60 \mathrm{GeV} / c^{2}$ for $\tau^{-} \rightarrow \pi^{-} \pi^{+} \pi^{-} \nu_{\tau}$, $0.50<m_{\pi^{-}} \pi^{0}<1.00 \mathrm{GeV} / c^{2}$ for $\left.\tau^{-} \rightarrow \pi^{-} \pi^{0} \nu_{\tau}\right)$.

We use detailed Monte Carlo simulations to determine for each $\tau$ decay channel the selection efficiencies $\varepsilon_{i}$ weighted by the corresponding branching fractions [3]. The systematic uncertainties in selection efficieny arise from tracking efficiency, neutral reconstruction, particle identification, and $\pi^{0}$ reconstruction. The total $B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}$ selection efficiency (see Table 【is ( $10.5 \pm 0.2$ )\%. Misreconstruction and contamination amongst the $\tau$-decay channels are taken into account.

Continuum and combinatorial $B$ background is determined by extrapolating the ARGUS function from the $m_{\mathrm{ES}}$ sideband into the $m_{\mathrm{ES}}$ signal region. The background that peaks in the $m_{\mathrm{ES}}$ signal region is determined from Monte Carlo simulations of $B^{+} B^{-}$events. Events where a $B^{0}$ is incorrectly reconstructed as a $B^{+}$provide a negligible contribution.

We correct the expected background, $b_{i}$, to take into account possible dependencies of the fitted ARGUS shape on a given discriminating variable ( $p_{\text {miss }}$, invariant masses, etc.). The correction factor is the ratio of the background expectations determined using two separate methods. In the first method, we estimate the background by scaling the number of events in the $m_{\mathrm{ES}}$ sideband using the ARGUS signal-to-sideband ratio. In the second method, we bin each discriminating variable and reweight the number of events, bin-by-bin, using the ARGUS signal-to-sideband ratio for each bin. The systematic error on $b_{i}$ is estimated as the deviation from unity of the total correction factor for each $\tau$-decay mode. The expected background and the total systematic uncertainty in each $\tau$-decay channel is reported in Table $\square$ along with the number $n_{i}$ of selected candidates in data.

The systematic uncertainty in $N_{B_{\text {had }}}(1.8 \%)$ is estimated as the change in the yield in the signal region in Fig. 2 when we use a double Gaussian as an alternative to the Crystal Ball function. Other models for the signal or the background distribution result in negligible changes.

We observe a total of $15 B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}$ candidates, which is consistent with the expected background of $17.2 \pm 2.1_{\text {stat }} \pm 1.3_{\text {syst }}$ events. The distribution of these events is also consistent with background.

We determine the $B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}$ branching fraction from the number of signal candidates $s_{i}$ expected for each $\tau$ decay mode, where $s_{i} \equiv N_{B_{\mathrm{had}}} \mathcal{B}\left(B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}\right) \varepsilon_{i}$. The results for each decay channel are combined using the estimator, $Q$. Here we define $Q$ to be $\mathcal{L}(s+b) / \mathcal{L}(b)$, where

$$
\begin{equation*}
\mathcal{L}(s+b) \equiv \prod_{i=1}^{n_{\mathrm{ch}}} \frac{e^{-\left(s_{i}+b_{i}\right)}\left(s_{i}+b_{i}\right)^{n_{i}}}{n_{i}!}, \mathcal{L}(b) \equiv \prod_{i=1}^{n_{\mathrm{ch}}} \frac{e^{-b_{i}} b_{i}^{n_{i}}}{n_{i}!} \tag{5}
\end{equation*}
$$

are the likelihood functions for signal-plus-background and background-only hypotheses and $n_{\text {ch }}$ is the total

TABLE I: Branching fraction ( $\mathcal{B}$ ) 3], efficiency $\left(\varepsilon_{i}\right)$, expected background $\left(b_{i}\right)$ with statistical and systematic errors, and observed data candidates ( $n_{i}$ ) for each reconstructed $\tau$ decay mode.

| selection | $\mathcal{B}(\%)$ | $\varepsilon_{i}(\%)$ | $b_{i}$ | $n_{i}$ |
| :--- | ---: | :---: | :---: | :---: |
| $e \nu \nu$ | $17.84 \pm 0.06$ | $3.4 \pm 0.1$ | $0.7 \pm 0.4 \pm 0.1$ | 2 |
| $\mu \nu \nu$ | $17.37 \pm 0.06$ | $1.9 \pm 0.1$ | $0.9 \pm 0.5 \pm 0.1$ | 0 |
| $\pi \nu$ | $11.06 \pm 0.11$ | $2.6 \pm 0.1$ | $1.3 \pm 0.6 \pm 0.2$ | 2 |
| $\pi^{-} \pi^{+} \pi^{-} \nu$ | $9.52 \pm 0.10$ | $0.6 \pm 0.1$ | $4.3 \pm 1.0 \pm 0.3$ | 4 |
| $\pi^{-} \pi^{0} \nu$ | $25.41 \pm 0.14$ | $2.0 \pm 0.1$ | $10.0 \pm 1.6 \pm 1.3$ | 7 |
| all | $81.20 \pm 0.22$ | $10.5 \pm 0.2$ | $17.2 \pm 2.1 \pm 1.3$ | 15 |

number of reconstructed $\tau$-decay channels.
Since we have no evidence of signal we set an upper limit on the branching fraction. The statistical and systematic uncertainties in the expected background are included in the estimator $Q$ by convolving the likelihood functions with a Gaussian distribution having as standard deviation the combined statistical and systematic errors in the background estimate [13]. We determine that $\mathcal{B}\left(B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}\right)<4.2 \times 10^{-4}$ (90\% C.L.).

To combine the results from the statistically independent hadronic and semileptonic samples, we first calculate the likelihood ratio estimator, $Q \equiv \mathcal{L}(s+b) / \mathcal{L}(b)$, using the likelihood functions from each method. We create a combined estimator from the product of the semileptonic $\left(Q_{\mathrm{sl}}\right)$ and hadronic $\left(Q_{\text {had }}\right)$ likelihood ratio estimators, $Q=Q_{\mathrm{sl}} \times Q_{\mathrm{had}}$. The measured branching fraction, which is the value that maximizes the likelihood ratio estimator, is $\left(2.3_{-1.3}^{+1.5}\right) \times 10^{-4}$. The lower one-standarddeviation bound does not include zero because of the small excess of signal events observed in the semileptonic analysis. Since this value is compatible with a zero branching fraction, we set a combined upper limit,

$$
\begin{equation*}
\mathcal{B}\left(B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}\right)<4.2 \times 10^{-4}(90 \% \text { C.L. }) \tag{6}
\end{equation*}
$$

The semileptonic analysis does not contribute significantly to the combined limit because of the observed small excess of signal events.

We use Eq. 1 Eq. 6 and the measured value of $\left|V_{u b}\right|$ to set a limit on $f_{B}$. We find $f_{B}<0.510 \mathrm{GeV}(90 \%$ C.L. $)$.

In conclusion, we have searched for $B^{-} \rightarrow \tau^{-} \overline{\nu_{\tau}}$ in the recoil of hadronic and semileptonic $B$ decays. We have set the most stringent upper limit to date on this process.

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