Jerome H. Friedman Stanford Linear Accelerator Center Stanford, California<br>Werner Stuetzle Department of Statistics Stanford University and Stanford Linear Accelerator Center Stanford, California

## I. INTRODUCTION


#### Abstract

Multivariate analysis can be thought of as methodology for detection, description and validation of structure in pdimensional ( $p>1$ ) point clouds. Classical multivariate analysis relies on the assumption that the observations forming the point cloud(s) have a Gaussian distribution. All information about structure is then contained in the means and covariance matrices, and the well-known apparatus for estimation and inference in parametric families can be brought to bear. The uncomfortable ingredient in this approach is the Gaussianity assumption. The data may be Gaussian with occasional outliers or even the bulk of the data simply might not conform to a Gaussian distribution. The first case is the subject of robust statistics and is not treated here. We discuss methods that do not involve any distributional assumptions. In this case,


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structure cannot be perceived by looking at a set of estimated parameters. An obvious remedy is to look at the data themselves, at the p-dimensional point cloud(s), and to base the description of structure on those views. As perception in more than three dimensions is difficult, the dimensionality of the data first has to be reduced, most simply by projection. Projection of the data generally implies loss of information. As a consequence, multivariate structure does not usually show up in all projections, and no single projection might contain all the information. These points are further illustrated in Chapter 2. It is therefore important to fudiciously choose the set of projections on which the model of the structure is to be based. This is the goal of projection pursuit procedures. A paradigm for multivariate analysis based on these ideas is presented in Chapter 3.

By design, projection pursuit methods are ideally suited for implementation or interactive computer graphics systems. The potential of interaction between user and algorithm was convincingly demonstrated in the PRIM-9 system for detection of hypersurfaces and clustering (see Fisherkeller et al [1974]); this system is discussed in Chapter 4. Procedures for multiple regression and multivariate density estimation based on projection pursuit are outined in Chapters 5 and 6 . Common properties of all projection pursuit procedures are discussed in Chapter 7.


> 2. DETECTION AND DESCRIPTION OF STRUCTURE WITH PROJECTIONS

Our goal is to detect and describe multivariate structure using projections of the data. However, structure, if present,


Fig. 1 Structured point cloud in two dimensions may not be apparent in all projections. This is illustrated by the following examples. Figure 1 shows a point. sample drawn from a bivariate distribution. The apparent structure of the point cloud (separation into two clusters) would be revealed by projection onto the subspace spanned by the vector (1, -1), whereas no structure would be apparent in a projection on the subspace spanned by the vector (1, 1).

The data for Figure 2 are generated from the regression model $Y=X_{1}+X_{2}+\epsilon$ with $\left(X_{1}, X_{2}\right)$ uniformly distributed in $[-1,1] \times[-1,1]$ and $\in \sim N(0,0.01)$. Figure 2 a shows a projection on the two-dimensional subspace spanned by $Y$ and the linear combination $Z=X_{1}+X_{2}$. This projection clearly shows the association between the predictors $X_{1}$ and $X_{2}$ and the response $Y$. A similar plot with $Z=X_{1}-X_{2}$, Figure $2 b$, is clearly less structured.


Fig. 2. (a) Projection of data from model $Y=X_{1}+X_{2}+\in$ on plane spanned by $Y$ and $Z=X_{1}+X_{2}$. ( $Y$ is plotted on the vertical axis). (b) Projection of data from model $Y=X_{1}+X_{2}+\in$ on plane spanned by $Y$ and $Z=X_{1}-X_{2}$.

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    These examplea show that it is important to search for
structured projections. This process is called projection
pursuit.
    It is easy to envision situations where not all the infor-
mation about the structure is contained in a single projection.
Consider the regression example above but with Y = X 
Figures 3a and 3b show two projections with Z Za}=\mp@subsup{X}{1}{}-\mp@subsup{X}{2}{}\mathrm{ and
Z
ple coordinate transformation }\mp@subsup{Z}{a}{}=\mp@subsup{X}{1}{}+\mp@subsup{X}{2}{},\mp@subsup{Z}{b}{\prime}=\mp@subsup{X}{1}{}-\mp@subsup{X}{2}{
allows one to express the response as Y = . 25 ( Za
is also interesting to notice that the quadratic dependence
on Za is washed out due to variability caused by the depend-
ence on }\mp@subsup{Z}{b}{}\mathrm{ , and vice versa. This suggests that once a struc-
tured projection has been found, the structure should be re-
moved so that one obtains a clearer view of what has not yet
been uncovered.
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Fig. 3. (a) Projection of data from model $Y=X_{1} \cdot X_{2}+\epsilon$ on plane spanned by $Y$ and $Z_{a}=X_{1}+X_{2}$. (b) Projection of data from model $Y=X_{1} . X_{2}+\epsilon$ on plane spanned by $Y$ and $Z_{b}=X_{1}-X_{2}$.

## 3. A PROJECTION PURSUIT PARADIGM

The discussion in the previous section motivates the following schema for a class of procedures modeling structure in multivariate data:
(1) Choose an initial model.

Repeat
(1i) Find a projection that shows deviation of the data from the current model, indicating previously undetected structure (Projection Pursuit).
(iii) Change the model to incorporate the structure found in (ii) (Model Update).

Until the current model agrees with the data in all projections.
-Such projection pursuit procedures can be implemented in batch mode. In this case, a figure of merit must be defined, which measures the amount of deviation between model and data revealed in a projection. This figure of merit usually is optimized by numerical search, although in some simple cases optimization can be done analytically. If the optimum figure of merit is less than a threshold, data and model are said to agree. Batch implementations of projection pursuit regression and density estimation are described in Sections 5 and 6.

By construction, profection pursuit procedures are ideally suited for implementation on interactive computer graphics systems. Interaction between program and user can help in

- search for interesting projections
- specification of model update
- termination
- interpretation of structure.


#### Abstract

Although projection pursuit procedures are useful in batch mode, their full power comes to bear in an interactive environment.


## 4. THE PRIM-9 SYSTEM

PRIM-9 (Fisherkeller et al, [1974]) is a system for visual inspection of up to nine-dimensional data, mainly intended for detecting clusters and hypersurfaces. It was implemented on an interactive computer graphics system which allows the modification of pictures in real time and thus makes it possible to generate movie-like effects. Its basic set of operations consists of

Projection: The observations can be projected on a subspace spanned by any pair of the coordinates; the projection is shown on a CRT screen.

Rotation: A subspace spanned by any two of the coordinates can be rotated. If the projection subspace and the rotation subspace share a common coordinate, the rotational motion causes the user to perceive a spatial picture of the data as projected on the three-dimensional subspace defined by the coordinates involved. When the user terminates rotation in a particular plane, the old coordinates in that plane are replaced by the current (rotated) coordinates. This makes it possible to look at completely arbitrary projections of the data, not necessarily tied to the original coordinates.

Masking: Subregions of the p-dimensional observation space can be specified, and only points inside the subregion are displayed. Under rotation, points willenter and leave the masked region.

Isolation: Points that are masked out (i.e., not visible)
can be removed, thus splitting the data into two subsets.
The first two operations, projection and rotation, allow the user to perform what one might call "manual projection pursuit". Isolation, the splitting of the data set into subsets, provides a rudimentary form of structure removal. When clustering is detected, the clusters can be separated and each of them examined individually. This process can be 1terated.

Although several have been implemented (Stuetzle \& Thoma [1978], Donoho et al [1981]), systems like PRIM-9 have not yet found widespread use. The main reason has been the price of the necessary computing equipment. The processing power needed to compute rotations at a reasonable update rate is quite high (on the order of 60000 multiplications per second for 1000 observations and 10 updates of the picture per second). Another major cost has been the graphics device, which must have a sufficiently high bandwidth (typically a megabaud). The situation, however, is rapidly changing. New 16-bit microprocessors provide a speed close to that required for an interactive use of projection pursuit procedures. The price of graphics systems, especially raster scan devices, is falling dramatically. The graphics system at SLAC used for the implementation of PRIM-9 cost $\$ 175,000$ in 1967. Today the price of a comparable system is $\$ 15,000$.
5. PROJECTION PURSUIT REGRESSION (PPR)

The goal of regression analysis is to find and describe the association between a response variable $Y$ and predictor variables $X_{1} \ldots X_{p}$, using a sample $\left\{\left(y_{i}, \underline{x}_{i}\right)\right\}_{i=1}^{n}$. PPR attempts
to construct a model for this association (or, in more classical terms, to estimate $E(Y \mid \underline{x})$ ) from the information contained in projections of the data on two-dimensional subspaces spanned by $Y$ and a linear combination $Z=\underline{\alpha}$. $\underline{X}$. The algorithm exactly follows our profection pursuit paradigm:
(1) Choose an initial model, for example $m_{0}(\underline{X})=$ const.

Repeat
(ii) Find a projection that shows deviation of the data from the model, i.e., find a direction such that the current residuals, $r_{i}=y_{i}-m\left(\underline{x}_{i}\right)$, show a dependence on $Z=\underline{\alpha}$. $\underline{X}$
(iii) Describe this dependence by a smooth function $s(Z)$. Update the model: $m(\underline{X}) \leftarrow m(\underline{X})+s(\underline{\alpha} \cdot \underline{X})$

Until data and model agree in all projections. ..
The model after $M$ iterations has the form

$$
\begin{equation*}
m(\underline{X})=m_{0}(\underline{X})+{\underset{m=1}{M} s_{m}\left(\underline{\alpha}_{m} \cdot \underline{X}\right) . ~ . ~ . ~}_{m} \tag{1}
\end{equation*}
$$

PPR allows the modeling of smooth but otherwise completely general regression surfaces. So far, a batch version has been implemented. Such an implementation requires the specification of figure of merit for projections and a method for summarizing a smooth dependence ("smoother"). Smoothing is generally accomplished by local averaging; the value of the smooth $s$ at a particular point $z$ is obtained by averaging the current residuals $r_{i}$ for those observations with values of $z_{i}$ close to $z$. The size of the neighborhoods within which averaging takes place is called the bandwidth of the smoother. A smoother suitable for use with PPR and guidelines for choor ing the bandwidth are described and discussed in Friedman and Stuetzle (1981).

A choice for the figure of merit is suggested by figures 2 a and 2 b . The (inverse) figure of merit is taken to be the residual sum of squares around the smooth of the current residuals versus $\underline{\alpha}$. $\underline{X}$. It is small in Figure $2 a$, where the smooth could closely follow the observations, and large in Figure $2 b$, where the smooth would be roughly constant. This definition of the figure of merit implies that in each iteration the model is updated along the direction for which the update yields the biggest reduction in residual sum of squares.

As with any stepwise procedure, one needs a criterion for stopping the iteration. Stopping too soon can increase the bias of the estimate, while not stopping soon enough can unduly increase its variance. "Optimal" termination of stepwise procedures has been studies (see Stone, [1981]); these methods can be applied here. In practice, the iteration is usually terminated subjectively, based on differences between successive values of the residual sum of squares. In addition, graphical inspection of $s_{m}(\underline{\alpha}, \underline{x})$ can be used to judge whether the corresponding term should be included in the model. If the graph of $s_{m}$ shows a noisy pattern with no systematic tendency, then its inclusion can only increase the variability of the estimate. On the other hand, a definite dependence indicates that $s_{m}$ deals with an inadequacy of the present model.

The following example illustrates the operation of PPR. A sample of 200 observations was generated according to the model
$Y=10 \sin \left(\pi X_{1} x_{2}\right)+20\left(x_{3}-0.5\right)^{2}+10 x_{4}+5 x_{5}+0 x_{6}+\epsilon$
with $\left(X_{1}, \ldots, X_{6}\right)$ uniformly distributed in $[-1,1]^{6}$ and $\in \sim N(0,1)$. Figure 4 a shows $Y$ plotted against the best single predictor,
$X_{4}$, and the corresponding smooth. (The response $Y$ is plotted on the vertical axis, $X_{4}$ on the horizontal axis. The "+" symbols represent data points, numbers indicated more than 1 data point. The smooth is represented by the "*" symbols.)
 found in the first iteration with $\underline{\alpha}_{1}=(0.41,0.51,-0.04$, $0.69,0.31,0.0)$. The association is seen to be approximately linear. The model after the first iteration thus is a plane which, in this case, closely coincides with the least squares plane through the data. Figure $4 c$ shows the residuals from this model ploted against the second linear combination $\underline{\alpha}_{2} \cdot \underline{X}$ found by the algorithm, with $\frac{\alpha}{-2}=(-0.14,0.0,0.99,0.04$, $0.0,-0.03)$. This iteration is seen to incorporate the quadratic dependence of the response on $X_{3}$ into the model. Figure $4 d$ shows the residuals after two iterations plotted against the third linear combination with $\underline{\alpha}_{3}=(.0 .70,0.72$, $0.01,0.03,0.02,0.00)$. Figure $4 e$ shows the residuals after three iterations ploted against the fourth linear combination, with $\underline{\alpha}_{4}=(0.80,-0.59,-0.10,0.04,0.01,0.0)$. The last two iterations are seen to model the interaction term $\sin \left(\pi X_{1} X_{2}\right)$. A further iteration failed to substantial ly improve the model.

For a more complete discussion of $P P R$ and additional examples, see Friedman and Stuetzle [1981].


Fig. 4. (a)Scatterplot of Y vs. $X_{4}$. (b) Scatterplot of $Y$ vs. $\underline{\alpha}_{1}$. X. (c) Scatterplot of $Y-s_{1}\left(\underline{\alpha}_{1} \cdot \underline{X}\right) \quad$ vs. $\underline{\alpha}_{2} \cdot \underline{X}$. (d) Scatterplot of $Y-s_{1}\left(\underline{\alpha}_{1} \cdot \underline{X}\right)-s_{2}\left(\underline{\alpha}_{2} \cdot \underline{X}\right)$ vs. $\underline{\alpha}_{3} \cdot \underline{X}$. (e) Scatterplot of $Y-s_{1}\left(\underline{\alpha}_{1} \cdot \underline{X}\right)-s_{2}\left(\underline{\alpha}_{2} \cdot \underline{X}\right)-s_{3}\left(\underline{\alpha}_{3} \cdot \underline{X}\right)$ vs. $\underline{\alpha}_{4} \cdot \underline{X}$.


Fig. $4 \mathrm{c}, \mathrm{d}$
(e)


Fig. 4 e
6. PROJECTION PURSUIT DENSITY ESTIMATION (PPDE)

The goal of density estimation is to estimate the multivariate distribution of a random vector $X$ on the basis of an i.i.d. sample $x_{1} \ldots x_{n}$. Our procedure again follows the projection pursuit paradigm:
(1) Choose an initial model for the density, for example, $m_{0}=m u l i v a r i a t e ~ n o r m a l$ with sample mean and covariance matrix.

## Repeat

(ii) Find a projection that shows deviation of the data from the model; i.c., find a direction such that $m(\underline{\alpha} \cdot \underline{X})$, the model marginal along $\underline{\alpha}$, differs from $p(\underline{\alpha} \cdot \underline{X})$, the (estimated) data marginal along $\underline{\alpha}$.
(iii) Define an "augmenting function" $f(\underline{\alpha} \cdot \underline{X})$ as the quotient of data and model marginals

$$
f(\alpha \cdot X)=\frac{p(\underline{\alpha} \cdot X)}{m(\underline{\alpha} \cdot \underline{X})}
$$

Update the model so that it and the data agree in the marginal along $\underline{\alpha}$ :

$$
m(\underline{X}) \leftarrow m(\underline{X}) \quad f(\underline{\alpha} \cdot \underline{X}) .
$$

Until data and model agree in all projections.
The model after $M$ steps of the iteration $1 s$ of the form

$$
\begin{equation*}
m(\underline{X})=m_{0}(\underline{X}) \cdot \prod_{m=1}^{M} f_{m}\left(\alpha_{m} \cdot \underline{X}\right) \tag{2}
\end{equation*}
$$

In step (iif) of the algorithm, the marginal of the data along $\alpha$ must be estimated and the marginal of the current model must be computed. The data marginal presents no problem. It can be estimated by projecting the data onto and using a one-dimensional kernel or near neighbor estimate. The analytic computation of the model marginal can be very difficult because it requires a (p-1)-dimensional integration We perform the integration by Monte Carlo, generating a sample from the model and proceeding as in the estimation of the data marginal.

As in the case of PPR, only a batch version of PPDE has so far been implemented. At each iteration, the direction $\underline{\alpha}$ is chosen such that the update of the current model yields the largest improvement in goodness-of-fit as measured by the likelihood of the sample. Termination rules are analogous to those used in PPR.

The following example illustrates the operation of PPDE. The data for the example are the concentration levels of four hormones in blood measurements of 256 children. The purpose of applying PPDE to these data is to determine if a Gausian distribution represents a reasonable approximation to the
data density. Figures 5a-5d compare the experimental data to a Monte Carlo sample drawn from a Gaussian density with the sample mean and covariance, as projected onto each of the measurement coordinates. The histogram of the experimental data is drawn with solid lines; the histogram of the Monte Carlo data is indicated by "*" symbols. Inspection of these projections indicates that although there are possibly some discrepancies, a Gaussian density might be a reasonable approximation to the data.

Figures 6a-6e show results for three iterations of PPDE. The solution direction $\underline{\sim}_{1}$ associated with the first iteration is mainly a combination of the second and third coordinate mea surements. The data distribution (Figure 6a) is seen to be somewhat skew and more peaked than the corresponding Gausian. The discrepancy between the data and the Gaussian model is much more pronounced in this projection than on any of the original coordinate measurements. Figure 6 b plots the augmenting function $f_{1}\left(\underline{\alpha}_{1} . \quad \underline{X}\right)$.

The second linear combination $\underline{\alpha}_{2}$ mainly involves the thtr and fourth coordinates. The principal difference between the current model $p_{1}(X)$ and the data is seen to be a substantial skewness to the left (Figure 6c). Figure 6d shows the corresponding augmenting function $f_{2}$. The linear combination associated with the third projection mostly involves the first and second coordinates. Although this iteration is trying to account for an apparent additional skewness of the data (Figure 6e), the effect is seen to be relatively small and perhaps not significant.

## DATA AND CURRENT MODEL PRDJこCTIDNS

(a)

data and current mødel prguectigns


Fig. 5. (a) Hormone data: Histogram of variable 1 with Gaussian Monte Carlo superimposed. (b) Hormone data: Histogram of variable 2 with Gaussian Monte Carlo superimposed. (c) Hormone data: Histograin of variable 3 with Gaussian Monte Carlo superimposed. (d) Hormone data: Histogram of variable 4 with Gaussian Monte Carlo superimposed.



Fig. $5 \mathrm{c}, \mathrm{d}$


Fig. 6. Hormone Data: (a) Histogram of lst solution linear combination $\alpha=(0.02,0.08,-0.59,0.14)$ with Gaussian model superimposed. (b) Augmenting function along lst solution inear combination $\alpha$ ( $0.02,0.80,-0.59,0.14$ ). (c) Histogramof 2 nd sosolution inear combination $\alpha_{2}=(-0.09,0.18,-0.45,-0.87)$ with current model Monte Carlo superimposed. (d) Augmenting function along second 1 inear combination $\chi_{2}=(-0.09,0.18,-0.45,-0.87)$. (e) Aistogram of third solution linear combination $\alpha=(0.45$, $0.88,0.16,-0.02$ ) with current model Monte Carlo superimposed.


Fig. $6 c, d$.


Fig. 6e.

Application of PPDE to these data reveals that a Gausian model provides a considerably less adequate description than indicated by the coordinate projections alone. The associated graphics gives some insight into the nature of the nonnormality of the data.

## 7. DISCUSSION

A11 projection pursuit procedures share some common advar tages:

- Since all estimation is carried out in a univariate setting the large bias of kernel or near neighbor estimates in high dimensions can often be avoided.
- PP procedures do not require specification of a metric in the observation space.
- Bias is encountered with stepwise procedures when many terms are required to provide a good representation of the model

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    underlying the data, but only a few can be estimated due to
    insufficient sample size. In these cases, it is imporeant
    that the first few terms be able to approximate a wide
    variety of functions so that the most salient features of
    the data can be modeled. In the limit M }->\infty\mathrm{ , any regres-
    sion function can be represented by (1), and any density
    can be represented by (2) (independent of the initial model),
    but even for moderate M, functions of those types constitute
    rich classes. In addition, the choice of the initial model
    permits the user to introduce any knowledge (s)he may have
    concerning the data, thereby allowing a further reduction
    in bias.
- As a data-analytic tool, projection pursuit procedures pro-
vide a set of directions { _ ...<M for exploring the differ-
ences between the initial model and the data. The fact that
at each stage the direction is chosen, for which the currert
model least adequately describes the data, makes them good
candidates for that purpose. A graphical comparison of the
projections of model and data, along with knowledge of the
initial model, can yield considerable insight into the mul-
tivariate data distribution. Pictorial representations of
each of the augmenting functions sm, respectively fm, along
with the particular directions over which they vary, can
also be quite informative since it is these functions that
actually comprise the model.
    There are situations in which projection pursuit proced-
ures can be expected not to perform well. Examples of regres-
sion functions requiring a large number of terms in equation
(1) are those with multiple peaks. Examples of unfavorable
density functions are those with highly concave isopleths or
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with sperically nested isopleths of the same de:sity value.
    In addition to regression and density estimation, the pro-
jection pursuit paradigm can be applied to the problems of
classification and robust estimation of covariance matrices.
All projection pursuit procedures use the same set of basic
operations, projection pursuit and model update. This should
allow the design of an interactive system for analysis of mul-
tivariate data that covers a wide range of problems and yet is
easy to learn and simple to operate.
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